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Default and the Maturity Structure in Sovereign Bonds *

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Abstract

During emerging market crises, government interest rate spreads rise, the debt maturity shortens and the spread on short-term bonds is higher than on long-term bonds. This paper studies the maturity composition of debt in a dynamic model with endogenous default, in which the price of debt compensates for the risk-adjusted losses from default. Short-term debt is better at inducing repayment because it does not require savings in the near future for repaying in the far future. Hence, short-term debt can raise more resources than long-term debt. However, issuing long-term debt provides a hedge against the need to roll-over short-term debt at high interest rate spreads. The trade-off between these two benefits is quantitatively important for understanding the maturity composition in emerging markets. When calibrated to data from Brazil, the model matches the dynamics in the maturity of debt issuances and its comovement with the level of spreads across maturities.

JEL Codes: F34, G15

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1 Introduction

Debt crises in emerging economies are often blamed on governments borrowing large amounts of short-term debt in international capital markets. Short-term borrowing leaves an economy with large amounts of debt to roll over, which becomes difficult when interest rates rise and access to external credit is restricted. This idea has motivated several recent empirical studies on debt and financial crises. For example, Rodrik and Velasco (2003) show, using evidence from a broad set of countries, that a high level of short-term foreign debt increases the likelihood of a crisis.¹ From this *ex post* point of view, it seems desirable to implement policies that would lengthen the maturity structure of emerging market economies' external liabilities. However, as documented by Broner, Lorenzoni, and Schmukler (2008), emerging market governments actively shift to shorter-maturity debt in a crisis and issue long-term debt in normal times. While this pattern may leave countries more exposed to roll-over crises in bad times, it suggests that there must be some benefits, *ex ante*, of shortening the maturity structure of debt precisely in such times. By the same token, the benefits of long-term debt must outweigh those of short-term debt in normal times.

In this paper, we develop a model of sovereign debt in which the borrower endogenously chooses a time-varying maturity structure of debt. The model accounts for the following observation in the data: when interest rate spreads rise, the short-term spread rises more than the long-term spread, while the maturity of newly issued debt shortens. Bond prices reflect the risk adjusted loss in case of default and are jointly determined along with the maturity structure of debt. Hence, we examine the maturity choice in a framework with realistic predictions for the term structure of interest rate spreads. Our model can rationalize shorter debt maturity during crises as the result of a liquidity advantage in short-term debt contracts; although these contracts carry higher spreads than longer term debt, they can deliver more resources to the country in times of high default risk.

We present data on prices and issuances of foreign-currency denominated bonds for four emerging market countries: Argentina, Brazil, Mexico, and Russia. We estimate *spread curves* – interest rate spreads over U.S. Treasury bonds across maturity – as well as the *duration* of bonds issued – a measure of the average time to maturity of payments on coupon-paying bonds. Across these four countries, within periods in which 2-year spreads are below their 25th percentile, the average duration of new debt is 7.1 years, and the average difference

¹Other detailed studies on individual cases include Calvo and Mendoza (1996) for Mexico; Radelet and Sachs (1998) for the East Asian economies; and Bevilacqua and Garcia (2002) for Brazil. Cole and Kehoe (1996) use a model of self-fulfilling crises to argue that the 1994 Mexican debt crisis could have been avoided if the maturity of government debt had been longer.

between the 10-year spread and the 2-year spread – the *slope* of the spread curve – is 2.3 percentage points. But when the 2-year spreads are above their 75th percentile, the average duration shortens to 5.7 years, while the slope of the spread curve is -0.5 percentage points. From this evidence we conclude that the maturity of debt shortens in times of high spreads and downward-sloping spread curves.

In our model, a borrower faces persistent income shocks and can issue long and short duration bonds. The borrower can default on debt at any point in time, but faces costs of doing so, in the form of lower income and exclusion from international financial markets. In equilibrium, default tends to occur in low-income, high-debt times, when the cost of debt payments outweighs the costs of default. Bond prices compensate for the expected loss from default as well as for risk premia.

Our model generates the dynamics of spread curves observed in the data because the endogenous probability of repayment is persistent, yet mean reverting, as a result of the dynamics of debt and income. In times of low debt and high income, default is unlikely in the near future, so spreads are low. Long-term spreads are higher than short-term spreads because default may become likely in the far future if the borrower receives a sequence of bad shocks and accumulates debt. Long-term spreads are also higher because of the risk premium on future changes in default probabilities. Conversely, in times of low income and high debt, default is likely in the near future, so spreads are high. Long-term spreads rise less than short-term spreads because the borrower’s likelihood of repaying may rise over a longer time horizon if a sequence of good shocks occurs and debt is reduced.

We calibrate the model to Brazil and find that it fits the observed dynamics of spreads well. When the spread on 2-year debt is below its 25th percentile, the 10-year spread is on average 2 percentage points higher than the 2-year spread, compared to a slope of 3 percentage points in Brazil. In periods when the 2-year spread is above its 75th percentile, the slope of the spread curve in the model inverts to -0.9 percentage points, compared to -1.5 percentage points in Brazil. We also use the model to decompose spreads into actuarially fair compensation for default and risk premia. We find that risk premia account for a substantial portion of the spread in periods of low spreads, but that the majority of the changes in spreads are accounted for by the dynamics of the default probability. For example, the risk premium accounts for one third of the 2-year spread when 2-year spreads are below their 25th percentile, but only for five percent when 2-year spreads are above their 75th percentile. In addition, the risk premium on long-term debt is always higher than on short-term debt, reflecting the fact that the risk premium is cumulated over a longer horizon. This allows the model to generate an average slope of the spread curve of about 1 percentage point, compared

to 1.5 percentage points in the Brazilian data.

The maturity structure of debt in the model reflects a trade-off between *liquidity* benefits of short-term debt and *hedging* benefits of long-term debt, both due to the presence of default. Short-term debt is a more liquid asset in the sense that consumption can be increased more effectively with short-term debt than with long-term debt. Issuing short-term debt induces the borrower to repay income in the near future to avoid the costs of defaulting. In contrast, the borrower cannot commit to saving sufficiently to repay long-term debt, so consumption cannot be raised as much with long-term debt.

Despite its liquidity benefits, short-term debt is risky, in the sense that rolling over a lot of short-term debt is expensive in future states with high interest rate spreads. Issuing long-term debt provides a hedge, because its value falls in bad states more than the value of short-term debt. This lowers the total debt burden in bad states relative to good states, effectively transferring resources across states in future periods. The value of outstanding long-term debt reflects cumulative risk-adjusted default probabilities over a longer horizon, so it is more sensitive to changes in default risk than the value of short-term debt.

The time-varying maturity structure in the model is a result of the time-varying valuation of the liquidity benefit of short-term debt and the hedging benefit of long-term debt. Periods of low default probabilities and upward-sloping spread curves correspond to states when the borrower is wealthy and values the hedging benefit that long-term debt provides. Thus, the portfolio is shifted toward long debt. Periods of high default probabilities and inverted spread curves correspond to states when the borrower is poor and credit is limited. These are times when liquidity is most valuable, and thus the portfolio is shifted toward shorter-term debt. In these periods, the average duration of debt is about 2.4 years shorter than in periods with low spreads, compared to a difference of about 1.7 years in Brazil. We can therefore rationalize higher short-term debt positions in times of crises as an optimal response to the illiquidity of long-term debt, and the tighter availability of its supply.

Risk premia in bond prices have substantial effects on default probabilities and the maturity of debt. In a counterfactual exercise, we find that raising risk premia makes borrowing more costly, which results in lower default probabilities in equilibrium. High risk premia also shorten the maturity of debt because they disproportionately raise the cost of issuing long-term debt.

Related Literature

Several recent papers study the maturity structure of sovereign debt. Jeanne (2009) argues that short-term debt gives incentives for sovereign governments to implement creditor-friendly

policies, because creditors can discipline the government by rolling over the debt only after desired policies are implemented. In our model, short-term debt also plays a role in inducing the government to repay when it lacks the ability to commit to repayment. Relative to Jeanne (2009), we consider why debt maturity shortens in times of high spreads. Broner, Lorenzoni, and Schmukler (2008) argue that emerging markets shift to short-term debt in a crisis because shocks to lenders' risk aversion raise the risk premium on long-term bonds more than on short-term bonds. In our model, we allow for time-varying risk premia, but the lenders' degree of risk aversion fluctuates with the borrower's income, instead of being driven by independent shocks.²

The mechanisms in our model build on two strands of the corporate finance literature. First, the basic tradeoff between issuing short-term and long-term debt in our model is similar to that discussed by Hart and Moore (1989, 1994). They develop a model in which an entrepreneur who cannot commit to repay seeks financing for a project from a lender.³ In addition, our results on the term structure of spreads mirror those derived in Merton (1974) for credit spread curves on defaultable corporate bonds. Our framework differs from Merton's in that the probability of default and the level and maturity composition of debt issuances are endogenous.

This paper is also related to the literature on the optimal maturity structure of government debt in closed economies. Angeletos (2002), Buera and Nicolini (2004) and Shin (2007) show that, when debt is not state contingent, a rich maturity structure of government bonds can be used to replicate the allocations obtained with state-contingent debt in economies with distortionary taxes as in Lucas and Stokey (1983). In these models, short- and long-term interest rate dynamics reflect the variation in the representative agent's marginal rate of substitution, which changes with the state of the economy. Lustig, Sleet, and Yeltekin (2006) show that higher interest rates on long-term debt relative to short-term debt reflect an insurance premium paid by the government for the benefits long-term debt provides in hedging against future shocks. Our paper shares with these papers the message that managing the maturity composition of debt can provide benefits to the government because of fluctuations in future interest rates.⁴

The model in this paper builds on the work of Aguiar and Gopinath (2006) and Arellano

²Bi (2007) and Niepelt (2009) present dynamic models focusing on other mechanisms. Bi shows how short-term debt reduces the value of outstanding long-term debt, while Niepelt focuses on smoothing borrowing costs across maturities.

³Diamond (1991) emphasizes a similar tradeoff in a model in which borrowers have private information over their future default risk.

⁴Neumeyer and Perri (2005) have shown that emerging markets face substantial fluctuations in interest rate spreads, which makes hedging interest rate fluctuations particularly useful.

(2008), who model equilibrium default with incomplete markets, as in the seminal paper on sovereign debt by Eaton and Gersovitz (1981). To connect our model to the data, we extend this framework to incorporate long debt of multiple maturities, rather than one-period debt as in these earlier papers, as well as a flexible pricing kernel that allows for risk premia in bond prices. In recent work, Chatterjee and Eyigungor (2009) and Hatchondo and Martinez (2009) show that models with a single long-term defaultable bond allow a better fit of emerging market data in terms of the volatility and mean of the country spread as well as debt levels. In addition, Borri and Verdelhan (2009) show that risk premia help to explain the variation of spreads across emerging markets with similar default probabilities.

The outline of the paper is as follows. Section 2 presents data on the dynamics of the spread curve and maturity composition for four emerging markets: Argentina, Brazil, Mexico, and Russia. Section 3 presents the theoretical model. Section 4 presents some examples to illustrate the mechanisms that determine the optimal debt portfolio. Section 5 presents all the quantitative results, and Section 6 concludes.

2 Emerging Markets Bond Data

We examine data on sovereign bonds issued in international financial markets by four emerging-market countries: Argentina, Brazil, Mexico, and Russia. We look at the behavior of the interest rate spreads over default-free bonds, across different maturities, and at the way the maturity of new debt issued covaries with spreads. We find that when spreads are low, governments issue long-term bonds more heavily and long-term spreads are higher than short-term spreads. When spreads rise, the maturity of bond issuances shortens and short-term spreads are higher than long-term spreads. Our findings confirm the earlier results of Broner, Lorenzoni, and Schmukler (2008), who showed in a sample of eight emerging economies that debt maturity shortens when spreads are very high.⁵

2.1 Spread Curves

We define the n -year *spread* for an emerging market country as the difference between the yield on a defaultable, zero-coupon bond maturing in n years issued by the country and on a zero-coupon bond of the same maturity with negligible default risk (for example, a U.S.

⁵Broner, Lorenzoni and Schmukler (2008) focus on the relationship between the term structure of *excess returns* and the average *maturity* of debt. In this section we construct measures of the term structure of *yield spreads* and the average *duration* of debt because these statistics provide the basis for the quantitative assessment of our model.

Treasury note). The spread is the implicit interest rate premium required by investors to be willing to purchase a defaultable bond of a given maturity. The *spread curve* depicts spreads as a function of maturity.

We denote the continuously compounded yield at date t on a zero-coupon bond issued by country i , maturing in n years, as $r_{t,i}^n$. The yield is related to the price $p_{t,i}^n$ of an n -year zero-coupon bond, with face value 1, through

$$p_{t,i}^n = \exp(-n \times r_{t,i}^n). \quad (1)$$

The n -year spread for country i at date t is given by: $s_{t,i}^n = r_{t,i}^n - r_{t,rf}^n$, where $r_{t,rf}^n$ is the yield of a n -year default-free bond.⁶

Since governments do not issue zero-coupon bonds in a wide range of maturities, we estimate a country's spread curve by using secondary market data on the prices at which coupon-bearing bonds trade. The estimation procedure consists of choosing a functional form for the spread curve to fit the discounted value of coupon payments to prices, following Svensson (1994) and Broner, Lorenzoni, and Schmukler (2008). We describe this procedure further in the Appendix and illustrate that the resulting pricing errors are small.

We compute spreads starting in March 1996 at the earliest and ending in May 2004 at the latest, depending on the availability of data for each country. Figure 1 displays the estimated spreads for 2-year and 10-year bonds for Argentina, Brazil, Mexico, and Russia.

Spreads are very volatile, and the difference between long-term and short-term spreads varies substantially over time. When spreads are low, long-term spreads are generally higher than short-term spreads. However, when the level of spreads rises, the gap between long and short-term spreads tends to narrow and sometimes reverses: the spread curve flattens or inverts. The time series in Figure 1 show sharp increases in interest rate spreads associated with Russia's default in 1998, Argentina's default in 2001, and Brazil's financial crisis in 2002.⁷ The expectation that the countries would default in these episodes is reflected in the high spreads charged on defaultable bonds.

To emphasize the pattern observed in the time series that short-term spreads tend to rise more than long-term spreads, in Figure 2 we display spread curves averaged across different

⁶Our data include bonds denominated in U.S. dollars and European currencies, so we take U.S. and Euro-area government bond yields as default-free.

⁷For Argentina and Russia, we do not report spreads after default on external debt, unless a restructuring agreement was largely completed at a later date. We use dates taken from Sturzenegger and Zettelmeyer (2005). For Argentina, we report spreads until the last week of December 2001, when the country defaulted. The restructuring agreement for external debt was not offered until 2005. For Russia, we report spreads until the second week of August 1998 and beginning again after August 2000 when 75% of external debt had been restructured.

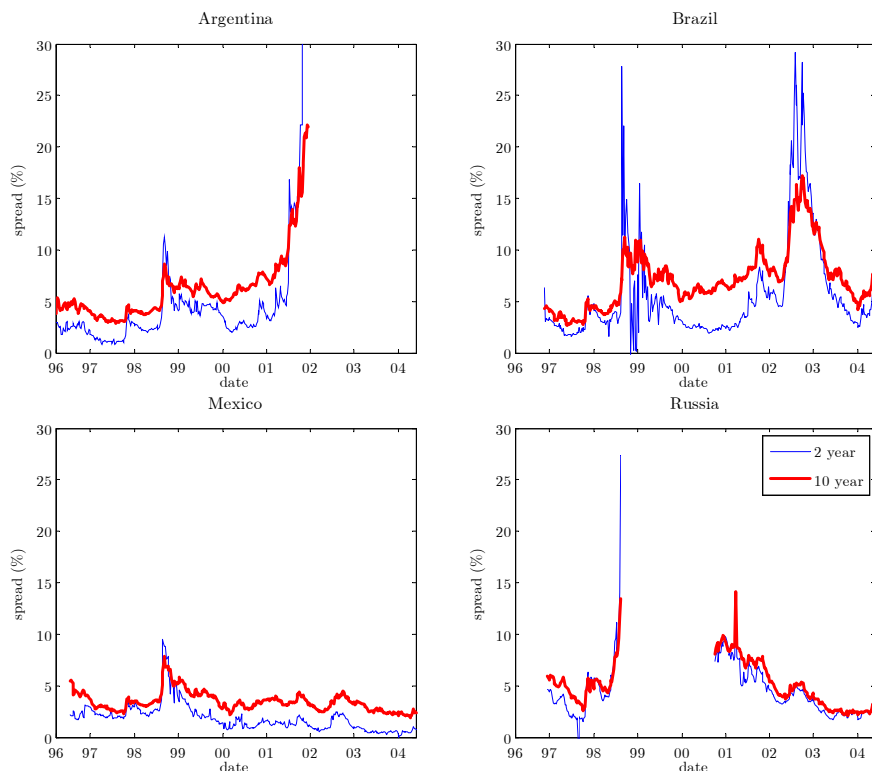


Figure 1: Time Series of Short and Long Spreads

time periods for each country: the overall average, the average within periods with the 2-year spread below its 10th percentile, and the average within periods with the 2-year spread above its 90th percentile. When spreads are low, the spread curve is upward sloping: long-term spreads are higher than short-term spreads. When spreads are high, short-term spreads rise more than long-term spreads. For Argentina, Brazil, and Russia, the spread curve becomes downward sloping in these times. For Mexico, which had relatively smaller increases in spreads during this time period, the spread curve flattens as short spreads rise more than long spreads.⁸

2.2 The Maturity Composition of Debt and Spreads

We now examine the maturity of new debt issued by the four emerging market economies during the sample period, and relate the changes in the maturity of debt to changes in

⁸The findings are similar to empirical findings on spread curves in corporate debt markets. Sarig and Warga (1989), for example, find that highly rated corporate bonds have low levels of spreads, and spread curves that are flat or upward-sloping, while low-grade corporate bonds have high levels of spreads, and average spread curves that are hump-shaped or downward-sloping.

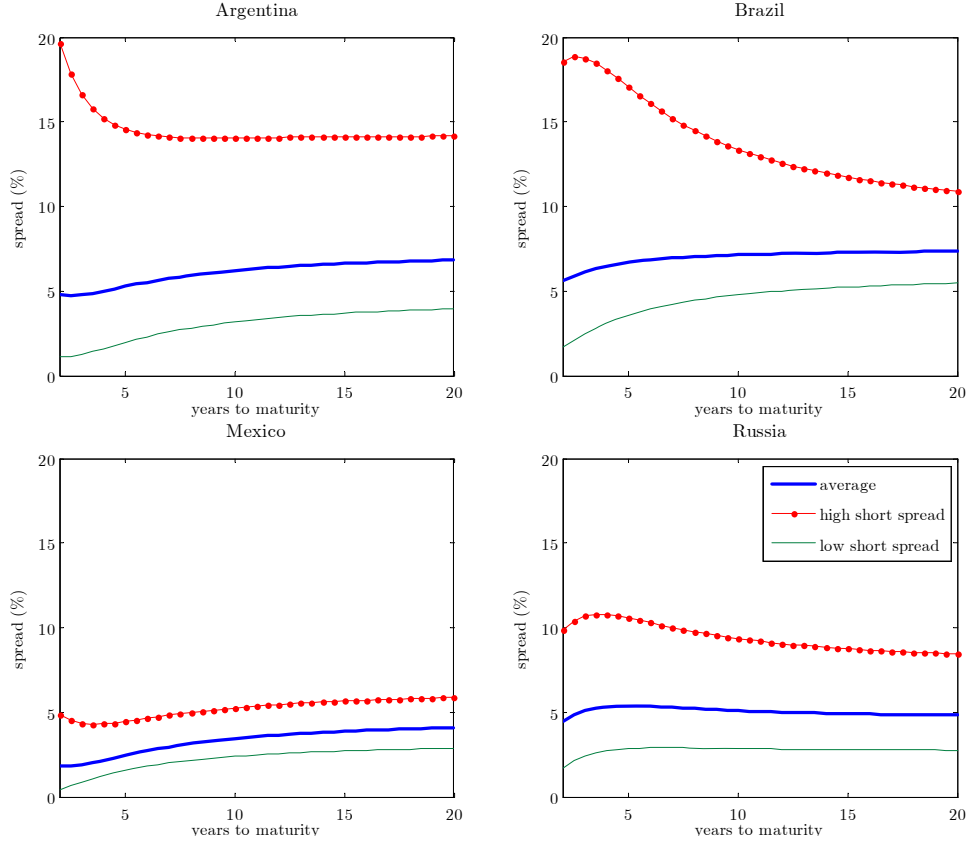


Figure 2: Spread Curves

spreads.⁹

In each week in the sample, we measure the maturity of debt as a quantity-weighted average maturity of bonds issued that week. We measure the maturity of a bond using two alternative statistics. The first is simply the number of years from the issue date until the maturity date. The second is the bond's *duration*, defined in Macaulay (1938) as a weighted average of the number of years until each of the bond's future payments. A bond issued at date t by country i , paying annual coupon c at dates n_1, n_2, \dots, n_J years into the future, and face value of 1 has duration $d_{t,i}(c)$ defined by

$$d_{t,i}(c) = \frac{1}{p_{t,i}(c)} \left(\sum_{j=1}^J \exp(-n_j r_{t,i}^{n_j}) n_j c + \exp(-n_J r_{t,i}^{n_J}) n_J \right), \quad (2)$$

⁹In addition to external bond debt, emerging countries also have debt obligations with multilateral institutions and foreign banks. However, marketable debt constitutes a large fraction of the external debt. The average marketable debt from 1996 to 2004 is 56% of total external debt in Argentina, 59% in Brazil, and 58% in Mexico (Cowan et al. 2006).

where $p_{t,i}(c)$ is the coupon bond's price, and $r_{t,i}^n$ is the zero-coupon yield curve. The time until each future payment is weighted by the discounted value of that payment relative to the price of the bond. A zero-coupon bond has duration equal to the number of years until its maturity date, but a coupon-paying bond has duration shorter than its time to maturity. We consider duration as a measure of maturity because it is more comparable across bonds with different coupon rates.

We calculate the average maturity and average duration of new bonds issued in each week by each country. Table 1 displays each country's averages of these weekly maturity and duration series within periods of high (above median) and low (below median) 2-year spreads.

Table 1: Average Maturity and Duration of New Debt				
2-year spread:	Maturity (years)		Duration (years)	
	< median	\geq median	< median	\geq median
Argentina	10.25	7.79	5.93	4.82
Brazil	13.66	7.06	6.37	4.73
Mexico	13.24	10.32	7.66	6.49
Russia	8.41	11.78	5.93	5.38

First, the table shows that duration tends to be much shorter than maturity. Because the yield on an emerging market bond is typically high, the principal payment at the maturity date is severely discounted, and much of the bond's value comes from coupon payments made before maturity.

Second, the average duration of debt is shorter when spreads are high than when they are low. Mexico, for example, issues debt that averages about 1.2 years longer in duration when the 2-year spread is below its median than when it is above its median. For all countries except Russia, this pattern also holds for the average time-to-maturity of bonds issued during periods of high spreads compared to low spreads: Mexico issues bonds that mature about 3 years sooner when spreads are high.

In Table 2, we report the results of several univariate panel regressions of duration on spread measures. We pool the data for the four countries and include country fixed effects. All coefficients are significant at the 1% level, and robust standard errors are reported in parentheses. Column I reports the effect of the 2-year spread on the duration of new issuances. The coefficient means that a one percentage point increase in the 2-year spread is associated

with a decrease in the duration of new debt issued by just under half a year. Column 2 shows a similar effect for the 10-year spread. These figures indicate that the covariation between the duration of new debt issuance and interest rate spreads is both economically large and statistically significant.¹⁰

Table 2: Regressions of Duration of New Issuances on Spreads

	I	II	III
Variable			
2-year spread	−0.446*** (0.116)		
10-year spread		−0.394*** (0.102)	
10/2 ratio			0.543*** (0.162)
R^2	0.201	0.195	0.157
No. of obs.	151	151	151

Note: All specifications include a constant. Robust standard errors are reported in parentheses, and *** denotes significance at the 1% level.

In Table 3, we emphasize the relationship between the spread curve *slopes* and average duration. The slope of the spread curve, defined here as the difference between the 10-year (long-term) and 2-year (short-term) spread, falls when the 2-year spread is high – the numbers in column 4 of Table 3 are smaller than those in column 3. During these times, however, the countries shift toward short-term debt, even though the spreads on long-term debt rise less than for short-term debt. In Brazil, for example, while the spread curve changes from depicting a 10-year spread that is about 3 percentage points above the 2-year spread to one that is 1.41 percentage points below the 2-year spread, the average duration of newly issued debt reduces by more than 2 years.

Column III of Table 2 shows that the duration of new debt is positively associated with how large the 10-year spread is relative to the 2-year spread. When the ratio of the long spread to the short spread increases by one, the duration of new debt rises by about half a year.

¹⁰These estimates mirror the findings in Broner, Lorenzoni, and Schmukler (2008). They show that a high spread level is a statistically significant determinant for a shorter maturity of debt issuances even after controlling for selection effects due to the fact that the timing of debt issuances is very irregular. Their empirical work treats the issuance of short-term or long-term debt as a discrete variable, whereas we use the continuous variable of duration as a measurement of maturity.

Table 3: Slope of Spread Curve and Average Duration of Issuances

	Duration (years)		Spread curve slope (%)	
			$s^{10} - s^2$	
short spread:	< 25th pct	≥ 75th pct	< 25th pct	≥ 75th pct
Argentina	6.33	5.02	2.16	-1.02
Brazil	6.94	4.63	3.09	-1.41
Mexico	8.79	6.75	1.98	0.79
Russia	6.46	6.12	0.78	0.22

2.3 Summary

The message of this section is that the spread curve and the maturity of bond issuances in emerging markets are time-varying. In particular, the level of spreads covaries negatively with the duration of new debt, and with the slope of the spread curve: when spreads are low, the slope of the spread curve is higher, and the maturity of new debt is longer, than when short-term spreads are high.

Standard asset pricing arguments would equate the price of a sovereign bond to the present discounted value of payments the bondholder expects to receive, adjusted for risk. Variation in interest rate spreads in this context, across maturity and over time, can come from a variety of sources, in particular, changes in the expected probability of default; changes in the amount lenders can recover in case of default; and variation in the lenders' compensation for risk, or risk premium. To understand the determinants of the maturity composition of debt, it is important to disentangle these three factors, since they may have different effects on the incentives to accumulate short- and long-term debt. Since we do not observe the time paths of the probability of default, the expected recovery rate, or the risk premium in the data, in the following sections we build a model in which the maturity structure of debt is endogenous, and these three factors play a role in pricing it.¹¹ When we calibrate this model to data for an emerging market economy, we can quantify the importance of these factors in explaining the relationship between spreads and debt maturity observed in the data.

¹¹An alternative approach to bond pricing with default is that of Duffie and Singleton (1999), used for sovereign debt in Duffie, Pedersen and Singleton (2003) and Pan and Singleton (2008). These papers back out risk- and recovery-adjusted probabilities of default from bond price data, but they need to specify an exogenous stochastic process for default events, and they do not consider the issuance of debt by the borrower.

3 The Model

We consider a dynamic model of defaultable debt that includes bonds of short and long duration. A small open economy receives a stochastic stream of income, y that follows a Markov process with compact support and transition function $f(y_t, y_{t+1})$. The economy trades two bonds of different duration with international lenders. Financial contracts are unenforceable, so the economy can default on its debt at any time. If the economy defaults, it temporarily loses access to international financial markets and also incurs direct costs.

The representative agent in the small open economy (henceforth, the “borrower”) receives utility from consumption c_t and has preferences given by

$$E \sum_{t=0}^{\infty} \beta^t u(c_t), \quad (3)$$

where $0 < \beta < 1$ is the time discount factor and $u(\cdot)$ is increasing and concave.

The borrower issues debt in the form of two types of perpetuity contracts with coupon payments that decay geometrically. We let $\{\delta_S, \delta_L\} \in [0, 1]$ denote the “decay factors” of the payments for the two bonds. A perpetuity with decay factor δ_m is a contract that specifies a price q_t^m and a loan face value ℓ_t^m such that the borrower receives $q_t^m \ell_t^m$ units of goods in period t and promises to pay, conditional on not defaulting, $\delta_m^{n-1} \ell_t^m$ units of goods in every future period $t + n$. The decay of each perpetuity is related to its duration: a bond of this type with rapidly declining payments has a larger proportion of its value paid early on, and therefore a shorter duration, than a bond with more slowly declining payments. We let $\delta_S < \delta_L$, so that δ_S is the decay of the perpetuity with short duration and δ_L is the decay of the perpetuity with long duration. We will refer to the perpetuities with decay factors δ_S and δ_L throughout as short and long bonds, respectively. Each one of our bonds resembles the long-duration bond in Hatchondo and Martinez (2009) and Rudebusch and Swanson (2008). By having two such assets, we allow the borrower to choose the maturity composition of debt each period.

At every date t the economy has outstanding all past bond issuances. Define b_t^m , the stock of bonds of duration m at time t , as the total payments due in period t on all past issuances of type m , conditional on not defaulting:

$$b_t^m = \sum_{j=1}^t \delta_m^{j-1} \ell_{t-j}^m = \ell_{t-1}^m + \delta_m \ell_{t-2}^m + \delta_m^2 \ell_{t-3}^m + \dots + \delta_m^t \ell_0^m + b_0^m,$$

where b_0^m is given. Thus, the accumulation for the stocks of short and long perpetuities can

be written recursively by the following laws of motion:

$$\begin{aligned} b_{t+1}^S &= \delta_S b_t^S + \ell_t^S \\ b_{t+1}^L &= \delta_L b_t^L + \ell_t^L \end{aligned} \quad (4)$$

With these definitions, we can compactly write the borrower's budget constraint conditional on not defaulting. Purchases of consumption are constrained by the endowment less payments on outstanding debt, $b_t^S + b_t^L$, plus the issues of new short bonds ℓ_t^S at price q_t^S and long bonds ℓ_t^L at a price q_t^L :

$$c_t = y_t - b_t^S - b_t^L + q_t^S \ell_t^S + q_t^L \ell_t^L \quad (5)$$

The borrower chooses new issues of perpetuities from a menu of contracts where prices q_t^S and q_t^L for are quoted for each pair (b_{t+1}^S, b_{t+1}^L) .

If the borrower defaults, we assume that all outstanding debts and assets ($b_t^S + b_t^L$) are erased from the budget constraint, and the economy cannot borrow or save, so that consumption equals output. In addition, the country incurs output costs:

$$y_t^{def} = \begin{cases} y_t & \text{if } y_t \leq (1 - \lambda)\bar{y} \\ (1 - \lambda)\bar{y} & \text{if } y_t > (1 - \lambda)\bar{y} \end{cases},$$

where \bar{y} is the mean level of output. This specification, following Arellano (2008), assumes that borrowers lose a fraction λ of output if output is above a threshold.

3.1 Recursive Problem

We represent the borrower's infinite horizon decision problem as a recursive dynamic programming problem. The model has two endogenous states – the stocks of each type of debt, b_t^S and b_t^L – and one exogenous state, the income of the economy, y_t . The state of the economy at date t is then given by $(b_S, b_L, y) \equiv (b_t^S, b_t^L, y_t)$.

At any given state, the value of the option to default is given by

$$v^o(b_S, b_L, y) = \max_{c,d} \{v^c(b_S, b_L, y), v^d(y)\}, \quad (6)$$

where $v^c(b_S, b_L, y)$ is the value associated with not defaulting and staying in the contract and $v^d(y)$ is the value associated with default.

Since we assume that default costs are incurred whenever the borrower fails to repay its

obligations in full, the model will generate default on all outstanding debt, both short and long term. When the borrower defaults, output falls to y_{def} , and the economy is temporarily in financial autarky; θ is the probability that it will regain access to international credit markets each period. The value of default is then given by the following:

$$v^d(y) = u(y_{def}) + \beta \int_{y'} [\theta v^o(0, 0, y') + (1 - \theta)v^d(y')] f(y, y') dy'. \quad (7)$$

We are taking a simple route to model both costs of default that seem empirically relevant: exclusion from financial markets and direct costs in output. Moreover, we assume that the default value does not depend on the maturity composition of debt prior to default. This captures the idea that the maturity composition of defaulted debt is not relevant for the restructuring procedures that allow the economy to reenter the credit market.¹²

When the borrower chooses to remain in the contract, the value is the following:

$$v^c = \max_{\{b'_S, b'_L, \ell_S, \ell_L, c\}} \left(u(c) + \beta \int_{y'} v^o(b'_S, b'_L, y') f(y, y') dy' \right) \quad (8)$$

subject to the budget constraint:

$$c - q_S(b'_S, b'_L, y) \ell_S - q_L(b'_S, b'_L, y) \ell_L = y - b_S - b_L \quad (9)$$

and to the laws of motion for the stock of perpetuities of short and long duration:

$$b'_S = \delta_S b_S + \ell_S$$

$$b'_L = \delta_L b_L + \ell_L.$$

The borrower decides on optimal debt levels b'_S and b'_L to maximize utility. The borrower takes as given that each contract $\{b'_S, b'_L\}$ comes with specific prices $\{q^S, q^L\}$ that are contingent on today's state y . The decision of whether to remain in the credit contract or default is a period-by-period decision, so that the expected value from next period forward in (8) incorporates the option to default in the future.

The default policy can be characterized by default sets and repayment sets. Let the repayment set, $R(b_S, b_L)$, be the set of output levels for which repayment is optimal when

¹²This is consistent with empirical evidence regarding actual restructuring processes, where the maturity composition of the new debt obligations is part of the restructuring agreement (Sturzenegger and Zettelmeyer 2005).

short- and long-term debt are (b_S, b_L) :

$$R(b_S, b_L) = \{y \in Y : v^c(b_S, b_L, y) \geq v^d(y)\}, \quad (10)$$

and let the complement, the default set $D(b_S, b_L)$, be the set of output levels for which default is optimal for debt positions (b_S, b_L) :

$$D(b_S, b_L) = \{y \in Y : v^c(b_S, b_L, y) < v^d(y)\}. \quad (11)$$

When the borrower does not default, optimal new debt takes the form of two decision rules mapping today's state into tomorrow's debt levels:

$$\begin{aligned} b'_S &= \tilde{b}_S(b_S, b_L, y) \\ b'_L &= \tilde{b}_L(b_S, b_L, y) \end{aligned} \quad (12)$$

Given this characterization of debt and default decisions, we can now define the equilibrium bond prices at which lenders are willing to offer contracts.

3.2 Bond Prices, Spreads, and Duration

Lenders are perfectly competitive, and value payoffs across states and time according to a pricing kernel, or stochastic discount factor, $M(y_t, y_{t+1})$, which we specify further below. Lenders are willing to purchase a defaultable bond at a price equal to the risk-adjusted expected discounted value of payments received from the bond.¹³

Each new issue of debt $\ell_t^S > 0$ or $\ell_t^L > 0$ by the borrower is a promise to pay a coupon payment every period in the future, conditional on not defaulting up to that period. If ℓ_t^S or ℓ_t^L is negative, then the borrower is repurchasing some of its debt. If the borrower defaults in period t , we assume that lenders recover a fraction $\varphi(b_t^S, b_t^L)$. The price of a new debt issue, then, is the discounted sum of the value of the promised coupon payments, adjusted by the cumulative probability of repayment, plus the value of recovery in case of default. If the borrower's state is (y_t, b_t^S, b_t^L) , the prices q_t^m for loans ℓ_t^m of duration $m = S, L$ are given

¹³Cochrane and Piazzesi (2005) show that risk premia are crucial for understanding government bond prices. Risk premia allow us to match bond prices and spreads better, and we consider their effects on the maturity structure of debt in the numerical results. We discuss in detail in Section 5 why we specify the pricing kernel as a function of the borrower's income.

by

$$q_t^m = \sum_{n=1}^{\infty} \int_{R_{t+1}} \cdots \int_{R_{t+n-1}} \Phi_{t+n}^m \tilde{f}(y_{t+n-2}, y_{t+n-1}) \cdots \tilde{f}(y_t, y_{t+1}) dy_{t+n-1} \cdots dy_{t+1} \quad (13)$$

where

$$\Phi_{t+n}^m = \delta_m^{n-1} \int_{R_{t+n}} \tilde{f}(y_{t+n-1}, y_{t+n}) dy_{t+n} + \varphi(b_{t+n}^S, b_{t+n}^L) \int_{D_{t+n}} \tilde{f}(y_{t+n-1}, y_{t+n}) dy_{t+n} \quad (14)$$

and $\tilde{f}(y_t, y_{t+1}) = M(y_t, y_{t+1}) f(y_t, y_{t+1})$.¹⁴ In (13)-(14), we use the shorthand $R_t = R(b_t^S, b_t^L)$, $D_t = D(b_t^S, b_t^L)$ for repayment and default sets. The term Φ_{t+n}^m is the risk-adjusted expected payoff of type m bonds in period $t+n$, conditional on the borrower having repaid up to period $t+n-1$. It is given by a weighted sum of the coupon payment, δ_m^{n-1} , in those states in which the borrower repays, and the recovery amount, $\varphi(b_{t+n}^S, b_{t+n}^L)$, in those states in which the borrower defaults.

Note that the price q_t^m of new debt issuances depends on current income, y_t , as it influences expectations of future output realizations which determine future default decisions. The price also depends on the entire future sequence of debt levels, $\{b_{t+n}^S, b_{t+n}^L\}_{n=1}^{\infty}$, since the outstanding debt in any period determines the decision to default, given the income level. However, we can transform the infinite sum in (13) into a recursive expression for q_t^m by assuming that the lender forecasts the future debt levels using the borrower's own decision rules for debt, defined in (12). The sum in (13) can then be written as

$$\begin{aligned} & \int_{R(b'_S, b'_L)} \tilde{f}(y, y') dy' + \int_{D(b'_S, b'_L)} \varphi(b'_S, b'_L) \tilde{f}(y, y') dy' \\ & + \int_{R(b'_S, b'_L)} \delta_m \left[\int_{R(\tilde{b}_S(b'_S, b'_L, y), \tilde{b}_L(b'_S, b'_L, y))} \tilde{f}(y, y') \tilde{f}(y', y'') dy'' \right] dy' + \dots \end{aligned}$$

where each future debt level is replaced in sequence by the optimal decision rules $\tilde{b}_S(b'_S, b'_L, y)$ and $\tilde{b}_L(b'_S, b'_L, y)$. Prices for debt then satisfy the functional equations:

$$\begin{aligned} q_S(b'_S, b'_L, y) &= \int_{R(b'_S, b'_L)} \left[1 + \delta_S q_S(\tilde{b}_S(b'_S, b'_L, y'), \tilde{b}_L(b'_S, b'_L, y'), y') \right] \tilde{f}(y', y) dy' \\ &+ \int_{D(b'_S, b'_L)} \varphi(b'_S, b'_L) \tilde{f}(y, y') dy' \end{aligned} \quad (15)$$

¹⁴ \tilde{f} is just a normalization of the “risk-neutral” probability density, $\frac{1}{E_t[M(y_t, y_{t+1})]} M(y_t, y_{t+1}) f(y_t, y_{t+1})$.

$$q_L(b'_S, b'_L, y) = \int_{R(b'_S, b'_L)} \left[1 + \delta_L q_L(\tilde{b}_S(b'_S, b'_L, y'), \tilde{b}_L(b'_S, b'_L, y'), y') \right] \tilde{f}(y', y) dy' \quad (16)$$

$$+ \int_{D(b'_S, b'_L)} \varphi(b'_S, b'_L) \tilde{f}(y, y') dy'$$

A recursive equilibrium for this economy is (i) a set of policy functions for consumption $\tilde{c}(b_S, b_L, y)$, new issuances for short-term debt $\tilde{\ell}_S(b_S, b_L, y)$ and long-term debt $\tilde{\ell}_L(b_S, b_L, y)$, perpetuity stocks for short-term debt $\tilde{b}_S(b_S, b_L, y)$ and long-term debt $\tilde{b}_L(b_S, b_L, y)$, repayment sets $R(b_S, b_L)$, and default sets $D(b_S, b_L)$, and (ii) price functions for short debt $q_S(b'_S, b'_L, y)$ and long debt $q_L(b'_S, b'_L, y)$, such that:

1. Taking as given the bond price functions $q_S(b'_S, b'_L, y)$ and $q_L(b'_S, b'_L, y)$, the policy functions $\tilde{b}_S(b_S, b_L, y)$, $\tilde{b}_L(b_S, b_L, y)$, $\tilde{\ell}_S(b_S, b_L, y)$, $\tilde{\ell}_L(b_S, b_L, y)$ and $\tilde{c}(b_S, b_L, y)$, repayment sets $R(b_S, b_L)$, and default sets $D(b_S, b_L)$ satisfy the borrower's optimization problem.
2. The bond price functions $q_S(b'_S, b'_L, y)$ and $q_L(b'_S, b'_L, y)$ satisfy equations (15) and (16).

We consider risk premia and recovery in bond prices through the functions M and φ because these features are important for quantitatively matching the behavior of spreads in the data. However, we take a reduced-form approach to these features, because our focus is on how the borrower chooses the maturity structure of debt, given the equilibrium price schedules q_S and q_L . From this perspective, risk compensation and recovery affect the borrowing decision only through their effects on the price schedules, so a more detailed model of international lenders should have the same effect as a sufficiently flexible reduced-form description of their behavior.

To compare the model's predictions to the data, we define the yield on each bond as in the data, as the implicit constant interest rate at which the discounted value of the bond's coupons equal its price. That is, given an equilibrium price q_m at any state, the yield r_m is defined from

$$q_m = \sum_{n=1}^{\infty} e^{-nr_m} \delta_m^{n-1}.$$

So,

$$r_S = \log \left(\frac{1}{q_S} + \delta_S \right) \text{ and } r_L = \log \left(\frac{1}{q_L} + \delta_L \right). \quad (17)$$

We define the spread as the difference between the yield on a defaultable bond and the default-free rate:

$$s_S = r_S - r_S^* \text{ and } s_L = r_L - r_L^*.$$

The default-free rates r_m^* are the analogues of (17) defined from the prices of default-free bonds.

As output and debt change, the probability of default varies over time, and therefore the prices of long-term and short-term debt differ, since they each put different weights on repayment probabilities in the future, as seen in (13). Spreads on short-term and long-term bonds therefore generally differ, and the relationship between the two spreads changes over time, so that the spread curve is time-varying.

Finally, we define the duration of debt issued at each date as the weighted average of the time until each coupon payment, with the weights determined by the fraction of the bond's value on each payment date:

$$d_m = \frac{1}{q_m} \sum_{n=1}^{\infty} n e^{-nr_m} \delta_m^{n-1}.$$

So,

$$d_S = \frac{1}{1 - \delta_S e^{-r_S}} \text{ and } d_L = \frac{1}{1 - \delta_L e^{-r_L}}. \quad (18)$$

4 Default and Optimal Maturity

In this section we illustrate the mechanisms that determine the maturity composition of debt in a simplified, three-period version of our model, in which the borrower receives income over time, but prefers to consume sooner rather than later. In this environment, debt allows the borrower to transfer income between states of the world, particularly from future states to the present, and the price of debt is determined by the borrower's willingness to repay. We show that short-term debt is more effective than long-term debt at transferring income from the near future to the present. However, long-term debt allows the borrower to avoid rolling over large short-term loans at unfavorable prices, which allows resources to be more efficiently transferred from high-income to low-income states. We use this example to derive statistics that are useful in measuring the relative benefits of short-term and long-term debt in our quantitative results in the next section.

The three periods are labelled $t = 0, 1, 2$. The borrower's preferences are linear over consumption in each period, given by

$$U = E[c_0 + \beta c_1 + \beta^2 c_2],$$

with $\beta < 1$, so that the borrower prefers to front-load consumption. In period 0, the borrower's income equals zero, and in periods 1 and 2 income y_1 and y_2 can take on one of two

values, y_H or y_L with $y_H > y_L$, according to the following process:

$$\begin{aligned}\Pr(y_1 = y_H) &= \alpha \\ \Pr(y_2 = y_H | y_1 = y_H) &= p_H \\ \Pr(y_2 = y_H | y_1 = y_L) &= p_L\end{aligned}$$

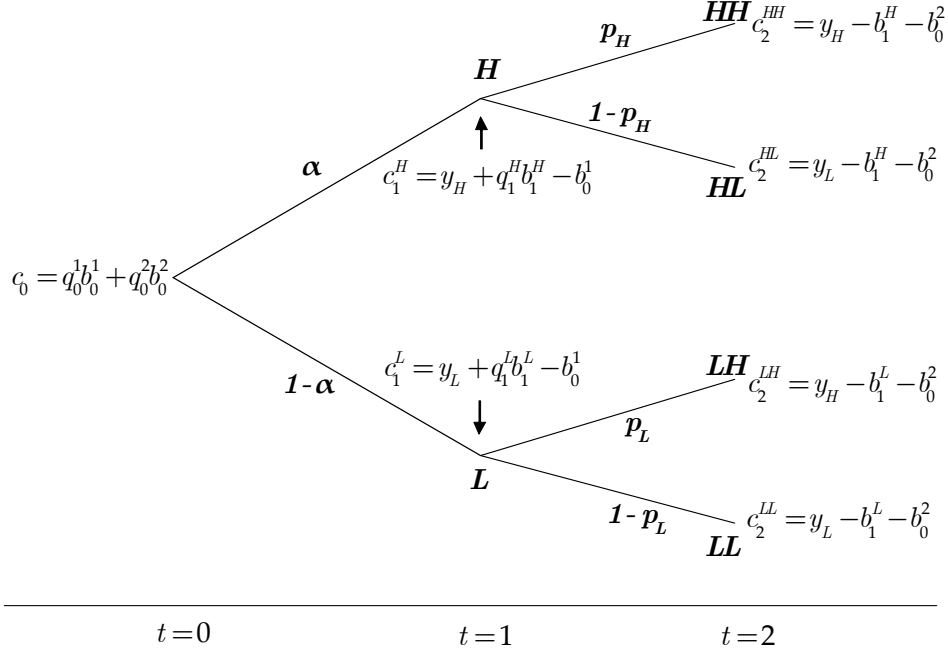


Figure 3: Structure of the Example

The income process and the borrower's consumption and borrowing possibilities are depicted in Figure 3. The borrower can issue one- and two-period zero coupon bonds b_0^1 and b_0^2 in period 0 at prices $q_0^1(b_0^1, b_0^2)$ and $q_0^2(b_0^1, b_0^2)$, and consumption is given by

$$c_0 = q_0^1(b_0^1, b_0^2)b_0^1 + q_0^2(b_0^1, b_0^2)b_0^2.$$

In period 1, conditional on not defaulting, the borrower can issue new one-period bonds b_1^i given the price schedule $q_1^i(b_1^i)$, where $i = L, H$ indexes the state, which corresponds to the realization of income y_1 . Consumption is equal to income plus new debt, minus repayment:

$$c_1^i = y_i + q_1^i(b_1^i)b_1^i - b_0^1,$$

In period 2, conditional on not defaulting, the borrower pays off long-term debt issued in

period 0 and short-term debt issued in period 1, and consumption equals income minus the repayment:

$$c_2^{ij} = y_j - (b_1^i + b_0^2) ,$$

where c_2^{ij} refers to consumption in the state with income history $y_1 = y_i, y_2 = y_j$, for $i, j = L, H$. We refer to states L and H in period 1 and states LL, LH, HL , and HH in period 2.

If the borrower defaults at any time, all outstanding debt is erased and consumption from then on is equal to $y_{def} = 0$. We assume that if the borrower is indifferent between defaulting and repaying, then he repays. Lenders are risk-neutral and do not discount the future, and they do not recover anything in case of default, i.e. the pricing kernel $M = 1$ and the recovery rate $\varphi = 0$, so the price of a loan is equal to the probability of repayment.

We assume throughout this section that parameters satisfy:

$$\frac{p_H y_H - y_L}{p_H (y_H - y_L)} < \beta < \left(\frac{p_H y_H - y_L}{p_H (y_H - y_L)} \right)^{1/2} \quad (19)$$

This restriction ensures that the borrower is sufficiently impatient so as to borrow enough that default happens with positive probability, but also patient enough so as to value the future resources lost from defaulting. Under this assumption, the optimal, time-consistent policy for the borrower is to default in period 2, only after two low shocks, i.e. in state LL , and to repay in all other states. Bond prices on loans that are consistent with this default policy are:

$$\begin{aligned} q_0^1 &= 1, q_0^2 = \alpha + (1 - \alpha) p_L \\ q_1^H &= 1, q_1^L = p_L \end{aligned} \quad (20)$$

In period 0, the borrower uses long-term debt to transfer income y_L to period 0, $b_0^2 = y_L$. This is because any two-period loan above y_L will not be repaid in state HL , violating the default policy. This leaves the borrower in period 1 with the ability to borrow additional income, up to what is consistent with the default policy, from period 2, $b_1^H = y_L - b_0^2 = 0$ and $b_1^L = y_H - b_0^2 = y_H - y_L$. This implies that in period 0, the borrower is able to issue one-period bonds $b_0^1 = y_L + p_L (y_H - y_L)$. The consumption allocation this implies is then

$$\begin{aligned} c_0 &= y_L + p_L (y_H - y_L) + (\alpha + (1 - \alpha) p_L) y_L \\ c_1^H &= (y_H - y_L) (1 - p_L), c_1^L = 0 \\ c_2^{HH} &= y_H - y_L, c_2^{HL} = 0 \\ c_2^{LH} &= 0, c_2^{LL} = y_{def} = 0 \end{aligned} \quad (21)$$

To evaluate the role each maturity of debt plays in getting to this consumption allocation, we consider what happens if either short-term or long-term debt is reduced.

Short-term debt induces repayment in the near future To illustrate the benefit of short-term debt, we consider what would happen if the borrower substituted long-term debt for short-term debt in period 0. Increasing long-term debt b_0^2 by *any* amount above y_L requires either defaulting in state HL as well as LL , or requires saving in period 1 to repay the higher debt in period 2. However, saving is not optimal from the perspective of period 1, so the borrower cannot commit to saving. In equilibrium, debt issued in either state in period 1 is nonnegative. The higher default probability lowers the price lenders are willing to pay for long-term bonds to $q_0^2 = \alpha p_H + (1 - \alpha) p_L$ from the value in (20), $\alpha + (1 - \alpha) p_L$. Furthermore, increasing b_0^2 above y_H is infeasible, since then q_0^2 would drop to zero, because of the inability to commit to saving in period 1. This means that the borrower is unable to transfer income from either state in period 1 to period 0, because long-term borrowing is limited by period-2 income.

In contrast, with short-term debt, the borrower is required to repay the income y_1 in the first period before borrowing against period 2 income. Since the borrower would value the resources lost from defaulting in period 1, short-term debt is a way to use the threat of punishment to enforce repayment in the short-term.

Long-term debt provides hedging Now we consider what would happen if the borrower tried to substitute short-term debt for long-term debt in the original allocation. This is infeasible because the borrower is unable to rollover any more short-term debt in period 1 state L . Under the optimal allocation derived above, consumption at all the nodes from period 1 state L onward – c_1^L , c_2^{LH} , and c_2^{LL} in (21) – are zero, so any additional debt due in period 1 state L cannot be repaid.

Long-term debt helps the borrower in this situation because it provides a hedge against the uncertainty in period 1. To see how this works, consider the resources available for consumption in period 1. Given that $b_1^H = y_L - b_0^2$ and $b_1^L = y_H - b_0^2$, consumption in period 1 is given by:

$$\begin{aligned} c_1^H &= y_H - b_0^1 + (y_L - b_0^2) \\ c_1^L &= y_L - b_0^1 + p_L (y_H - b_0^2) \end{aligned}$$

Decreasing b_0^1 in favor of increasing b_0^2 reduces the present value of the borrower's debt obligations – how much debt needs to be rolled over – in the low-income state, but leaves

it unchanged in the high-income state. In this sense, issuing long-term debt is a hedge. In effect, issuing long-term debt transfers resources from state H to state L , since increasing b_0^2 lowers c_1^H at a one-to-one rate, while it lowers c_1^L at the rate $p_L < 1$. This raises the value of resources at the rate $(1 - p_L)$. Since the borrower prefers to consume early, this transfer of income out of state H in period 1 changes consumption relative to the allocation with only one-period bonds by raising consumption in period 0 (by the present value $\alpha(1 - p_L)y_L$).

Equilibrium consumption and default with only one maturity of debt The previous two subsections illustrated that short-term and long-term debt provide distinct benefits, in the sense that the borrower cannot costlessly substitute one for the other. For completeness, we provide here the optimal allocations with only one maturity of debt. With only long-term debt, the borrower defaults in the additional state (HL) in period 2. With only short-term debt, though, the borrower does not actually borrow enough to face a roll-over crisis and default in period 1 state L ; borrowing is simply constrained in period 0.

The allocations with only one maturity of debt are as follows:

Only long-term debt	Only short-term debt
$b_0^2 = y_H$	$b_0^1 = y_L + p_L y_H$
$b_1^L = 0, b_1^H = 0$	$b_1^L = y_H, b_1^H = y_L$
$c_0 = (\alpha p_H + (1 - \alpha)p_L)y_H$	$c_0 = y_L + p_L y_H$
$c_1^H = y_H, c_1^L = y_L$	$c_1^H = (1 - p_L)y_H, c_1^L = 0$
$c_2^{HH} = 0, c_2^{HL} = y_{def} = 0$	$c_2^{HH} = y_H - y_L, c_2^{HL} = 0$
$c_2^{LH} = 0, c_2^{LL} = y_{def} = 0$	$c_2^{LH} = 0, c_2^{LL} = y_{def} = 0$

It is straightforward to show that the borrower attains strictly lower utility under either of these allocations than under the consumption allocation in (21) with both short-term and long-term debt.

Summary

In a standard incomplete markets model with fluctuating output and without default, a borrower would find the portfolio of long and short debt indeterminate if the risk-free rate were constant across time; the two assets would have payoffs that make them equivalent. However, in our model, the risk of default makes the two assets distinct, resulting in a tradeoff between issuing short-term and long-term debt. This tradeoff in our example is similar to that analyzed in Hart and Moore (1989, 1994), in a contracting problem between an investor and an entrepreneur who controls a productive project. Hart and Moore argue that the

tension between short-term and long-term debt is between defaulting early or defaulting later, but with larger losses. In our example, this tension limits the capacities of short-term and long-term borrowing differently.

The capacity for borrowing using long-term debt is limited by the inability to commit to saving income in the near future to repay farther in the future. Hence short-term debt is a better instrument for borrowing against resources in the near future, because large short-term loans are available at higher prices than long-term loans. We refer to this benefit as a *liquidity* advantage of short-term debt. However, the borrowing capacity of short-term debt is limited by the risk of not being able to roll over large amounts of short-term debt in low-income states. Long-term debt alleviates this problem by enabling a transfer of resources from high-income states to low-income states. We call this attribute a *hedging* benefit of long-term debt, because the outstanding value of long-term debt falls in states of the world in which resources are scarce. In our examples above, this hedging translated into higher current consumption, because we assumed the borrower did not value smoothing consumption across states; however, if the borrower were risk averse, long-term debt would enable smoothing consumption across states through this hedging benefit. This hedging mechanism is essentially the same as the role of long-lived securities in allocating risk across periods in Kreps (1982). Angeletos (2002), Buera and Nicolini (2004), and Lustig, Sleet, and Yeltekin (2008) emphasize this mechanism in models of the optimal maturity structure of government debt with incomplete markets. The difference in our model is that bond prices vary across states – and hence provide hedging – due to the government’s inability to commit to repaying, even in the absence of variation in the lender’s marginal rate of substitution.

These liquidity and hedging properties shape the optimal maturity structure of debt when the borrower can default. The quantitative relevance of each of these forces depends on the specifics of preferences and the income process. Thus, in the next section we quantify these two effects by calibrating our general model to an actual emerging market economy. To measure the liquidity and hedging mechanisms in our quantitative results in the next section, we use two simple measures that follow from the examples in this section. The liquidity advantage of short-term debt is reflected in its higher price and the ability to increase current consumption more than what is possible with long-term debt. The hedging advantage of long-term debt can be captured in the higher volatility of the value of outstanding long-term debt relative to short-term debt. The example above illustrates an extreme version of this: the value of short-term debt does not vary at all. Since the value of long-term debt varies more, it provides a better hedge than short-term debt.

5 Quantitative Analysis

5.1 Parameterization

We solve the model numerically to evaluate its quantitative predictions regarding the dynamic behavior of the optimal maturity composition of debt and the spread curve in emerging markets. We calibrate an annual model to the Brazilian economy.¹⁵

The utility function of the borrower is $u(c) = \frac{c^{1-\sigma}}{1-\sigma}$. The borrower's risk aversion coefficient is set to 2, which is a common value used in real business cycle studies. The stochastic process for output is a log-normal AR(1) process, $\log(y_{t+1}) = \rho \log(y_t) + \varepsilon_{t+1}$ with $E[\varepsilon^2] = \eta^2$. We discretize the shocks into a six-state Markov chain using a quadrature-based procedure (Tauchen and Hussey, 1991). We use annual series of GDP growth for 1960–2004 taken from the World Development Indicators to calibrate the volatility of output. Due to the short sample, rather than estimating the autocorrelation coefficient we choose an autocorrelation coefficient for the output process of 0.9, which is in line with standard estimates for developed countries. The decay parameters of the short and long bonds, δ_S and δ_L , are set such that the default-free durations equal 2 and 10 years. We also choose the probability of reentering financial markets, θ , so that the average length of time in exclusion is 6 years, consistent with data presented in Benjamin and Wright (2009) on the median length of sovereign debt renegotiations.

The function $\varphi(b_S, b_L)$ in the pricing expressions (15)-(16) is a reduced-form representation of the fraction of debt recovered after default. We set this function to $\varphi(b_S, b_L) = \exp(-(q_S^* b_S + q_L^* b_L))$, with q_S^* and q_L^* equal to the default-free prices. This functional form is a convenient way to capture two features of the recovery rate: it lies between zero and one, and it declines with the quantity of debt issued. Intuitively, this is motivated by the idea that there is a fixed amount of surplus that lenders are able to extract from borrowers in the event of a default, so this amount declines as a fraction of debt as debt grows. Yue (2006) shows that, in a model in which the debt recovery rate is endogenously determined by a bargaining process between lenders and a defaulting borrower, the recovery rate is decreasing and convex in debt, much like the function φ .

The risk premium in our model comes from the interaction of the lenders' pricing kernel with default outcomes. We can rewrite equation (15) (and analogously, (16)) as a typical

¹⁵The algorithm used to solve the model extends that in Arellano (2009) to allow for two bonds. In addition, we follow Hatchondo, Martinez and Saprizza (2009) in updating the value function and the bond price functions concurrently.

asset-pricing equation:

$$q_S = E [M' x'_S]$$

where

$$x'_S = \begin{cases} 1 + \delta_S q'_S & \text{in states in which the borrower repays} \\ \varphi & \text{in states in which the borrower defaults} \end{cases}$$

Written this way, the price of a bond is composed of the lender's discounted expected payoff plus a risk premium term:

$$q_S = E [M'] E [x'_S] + cov (M', x'_S)$$

Payoffs vary across states due to default events as well as changes in the future probability of default, reflected in q'_S . To the extent that this variation is negatively correlated with the pricing kernel, investors are compensated for this risk through paying a lower price q_S for debt. We specify the pricing kernel as follows:

$$M(y_t, y_{t+1}) = \exp \left(-r - \alpha \varepsilon_{t+1} - \frac{1}{2} \alpha^2 \eta^2 \right) \quad (22)$$

where r and α are parameters, $\varepsilon_{t+1} = \log y_{t+1} - \rho \log y_t$ is the shock to the borrower's income, and η is its variance. The term r represents the risk-free interest rate; we set this at 4% annually, which equals the average annual yield of a two year U.S. bond from 1996 to 2004. The term α controls correlated the pricing kernel is with innovations ε_{t+1} to the borrower's income level. (The term $\frac{1}{2} \alpha^2 \eta^2$ is a normalization.)

This pricing kernel is a variation on the one-factor term structure models discussed in, for example, Backus, Foresi, and Telmer (1998). We define the pricing kernel as a function of only the borrower's income because it is a parsimonious way to model risk premia that vary with the probability of default. In our model, default decisions depend on the model's entire state, (y, b_S, b_L) , but our shortcut is valid for two reasons: first, default probabilities in the model are highly correlated with income. Second, in the absence of default, our model with this pricing kernel would generate a default-free interest rate that is constant and a flat default-free term structure. This is because, from equations (15) or (16), the price q_m^* of a

default-free m -duration bond is given by¹⁶:

$$\begin{aligned} q_m^* &= \int (1 + \delta_m q_m^*) M(y, y') f(y', y) dy' \\ &= \frac{e^{-r}}{1 - \delta_m e^{-r}} \end{aligned} \tag{23}$$

so the yield, defined as in equation (17), is constant and independent of the bond's duration:¹⁷

$$\begin{aligned} r_m^* &= \log \left(\frac{1}{q_m^*} + \delta_m \right) \\ &= r \end{aligned}$$

Even though α is a constant, the size of the risk premium varies over time because the default probability, which depends on y , b_S , and b_L , varies over time. The borrower's income is a convenient state variable, because the borrower tends to default when income is low. Since income is persistent, low current income signals high default probabilities, and hence low prices, in the future. Therefore, since payoffs x_m are positively correlated with the borrower's income, the pricing kernel M in equation (22) generates a positive risk premium if $\alpha > 0$.

We calibrate the parameter α as well as the output cost after default, λ , and the borrower's time preference parameter β , to match the following moments of Brazilian data: the average 2- and 10-year spreads and the volatility of the trade balance relative output.

Table 4 summarizes the parameter values, and Table 5 presents the calibrated moments as well as other statistics from the model. The model matches the calibrated moments well, although it underestimates a bit the mean 10-year spread. The model predicts that consumption is about as variable as output, and that consumption is negatively correlated with spreads and the fraction of debt that is short term: in bad times, consumption is low, spreads are relatively high, and debt is mostly short. The high volatility of consumption relative to output and the countercyclicality of spreads are well documented features of emerging markets. The model predicts that the average recovery in default is 44%, as compared to the

¹⁶Using the fact that M is lognormal,

$$\begin{aligned} E[M(y, y') | y] &= \exp \left(E[\log M(y, y') | y] + \frac{1}{2} \text{var}[\log M(y, y') | y] \right) \\ &= \exp(-r) \end{aligned}$$

¹⁷In practice, since we discretize the state space of our model, we also have to normalize the pricing kernel so that default-free yields are constant in the discretized environment.

average recovery rate of 60% reported in sovereign defaults reported by Benjamin and Wright (2009).

Table 4: Parameters Values

	Value	Target
Lenders' discount rate	$r = 4\%$	U.S. annual interest rate 4%
Borrower's risk aversion	$\sigma = 2$	Standard value
Perpetuity decay factors	$\delta_S = 0.52$ $\delta_L = 0.936$	Default-free durations of 2 and 10 years
Stochastic structure	$\rho = 0.9, \eta = 0.017$	Brazil output
Probability of reentry	$\theta = 0.17$	Benjamin and Wright (2009)
Calibrated parameters		
Lenders' pricing kernel	$\alpha = 1.3$	Brazil average 2-year spread Brazil average 10-year spread volatility of trade balance
Output after default	$\lambda = 0.22$	
Borrower's discount factor	$\beta = 0.932$	

Table 5: Model Statistics

	Model	Data
<i>Targeted Moments</i>		
mean s_S (percent)	5.5	5.6
mean s_L (percent)	6.4	7.1
std(trade balance)/std(y)	0.36	0.36
<i>Other Moments</i>		
mean recovery rate	0.44	0.60
std(c)/std(y)	1.01	1.10
corr(c , 2-year spread)	-0.42	
corr(c , 10-year spread)	-0.34	
corr($c, q_S \ell_S / (q_S \ell_S + q_L \ell_L)$)	-0.77	

5.2 Results

We simulate the model, and in the following subsections we report statistics on the dynamic behavior of spreads and the maturity composition of debt from the limiting distribution of debt holdings. We first show that the model matches the data in generating time-varying differences in the pricing of short- and long-term debt, due to movements in the probability of default and risk premia. The behavior of prices generates time-varying liquidity and hedging benefits of the two assets, which rationalizes the maturity composition observed in the data.

Prices and Spreads

Default in our model happens when the economy has a low level of wealth, either due to low income or high debt. Since income and debt are persistent, states with low income and high debt tend to have high spreads, as the future probabilities of default, and therefore risk premia, are high.

We now compare spread dynamics in the model to the data. The series for the data are Brazilian 2- and 10-year spreads and prices from Section 2. For this comparison, we organize the data into quantiles based on the level of the short spread. Table 6 presents average spreads for short and long debt as well as the risk premium component of the model's spreads across periods when short spreads are below their 25th and 50th percentiles and above their 50th and 75th percentiles.

Table 6: Spread Curves

s_S percentile	DATA		MODEL			
	s_S	s_L	s_S	s_L	rp_S	rp_L
< 25	2.2	5.3	0.6	2.6	0.2	0.4
< 50	2.7	5.4	1.2	3.4	0.3	0.5
\geq 50	8.5	8.9	9.7	9.5	0.5	0.7
\geq 75	12.3	10.8	14.9	14.0	0.7	0.9
Mean	5.6	7.1	5.5	6.4	0.4	0.6

The first two columns of Table 6 present the short and long spreads in the data, and the last four columns present the model's predictions. In the model, when default is unlikely, both spreads are low, and the spread curve is upward-sloping: when the short spread is below its 25th percentile, for example, the average short spread is 0.6%, and the average long spread is 2.6%. In contrast, when the probability of default is higher, both spreads rise, and the spread curve becomes downward-sloping: when the short spread is above the 75th percentile, the average short spread is 14.9%, and the average long spread is 14.0%. Compared to the data for Brazil, the model captures well the dynamics of spreads curves and in particular the slope of the spread curve associated with periods of high and low spreads. When spreads are below the 25th percentile, the slope is about 2 percentage points, compared to a slope of 3 percentage points in Brazil. In periods with spreads above the 75th percentile, the slope of the spread curve in the model inverts to -0.9 percentage points, compared to -1.5 percentage points in Brazil. The model also generate an average slope of the spread curve of about 1 percentage point, compared to 1.5 percentage points in the Brazilian data.

Our model provides a decomposition of interest rate spreads into two parts: the actuarially fair compensation for expected losses from default, and the risk premium. The actuarially fair price q_m^{AF} of a bond of duration m is defined by taking the bond pricing equations (15)-(16) and substituting just the probability density, f , for the product of the pricing kernel and probability density, \tilde{f} :

$$\begin{aligned} q_m^{AF}(b'_S, b'_L, y) &= \int_{R(b'_S, b'_L)} \left[1 + \delta_m q_m^{AF}(\tilde{b}_S(b'_S, b'_L, y'), \tilde{b}_L(b'_S, b'_L, y'), y') \right] f(y', y) dy' \\ &\quad + \int_{D(b'_S, b'_L)} \varphi(b'_S, b'_L) f(y, y') dy', \end{aligned}$$

Then, the actuarially fair yield is given by $r_m^{AF} = \log(1/q_m^{AF} + \delta_m)$, and the *spread risk premium* is defined as $rp_m = r_m - r_m^{AF}$. The fifth and sixth columns of Table 6 show the term structure of spread risk premia in the model. Risk premia are positive, because the borrower tends to default in states with low income, and the lender's pricing kernel M is negatively correlated with income. Also, risk premia are generally small according to this measure, averaging less than one percentage point. However, they make up a relatively large fraction of the interest rate spreads in good times, when the spreads are low. For example, the risk premium accounts for one third of the 2-year spread when 2-year spreads are below their 25th percentile, but only for five percent when 2-year spreads are above their 75th percentile. In addition, the risk premium on long-term debt is always higher than on short-term debt, reflecting the fact that the risk premium is cumulated over a longer horizon. In effect, our model says that risk premia do not need to increase a lot in bad times to account for the large increases in spreads, because the expected probability of default rises so much. However, this decomposition does not address the question of how the degree of required risk compensation affects debt and default choices, and in turn the equilibrium term structure of spreads. We examine this question in more detail in subsection 5.3.

Much of the literature on empirical asset pricing focuses on returns as a measure of risk premia. In the context of sovereign bonds, Broner, Lorenzoni, and Schmukler (2008) show that excess returns – returns relative to default-free bonds – on long-term bonds rise more than excess returns on short-term bonds during a crisis, i.e. when spreads rise. Computing excess returns from our model's simulated data generates this pattern as well. Specifically, we find that the average excess returns, defined as $(1 + \delta_m q'_m)/q_m - (1 + r)$, equal 0.3% and 0.4% for short and long-term bonds in periods when the short spread is below its median. These returns increase to 7.9% and 10.4% respectively when the short spread is above its median. In our model, however, average excess returns in a simulation also capture changes in default

probabilities, not just risk premia. In fact, under actuarially fair pricing, the magnitude and pattern of average excess returns are similar to those under pricing with risk premia. We choose not to focus on average excess returns, because in our model they are not zero in the absence of risk premia.

Underlying the time-varying spreads is the interaction of the dynamics of income and debt with the price schedules for short and long debt. (Figure 4, in the Appendix, illustrates the equilibrium price schedules for short debt $q_S(b'_S, b'_L, y)$ and long debt $q_L(b'_S, b'_L, y)$.) However, the mapping from discount prices to spreads is not linear (eq. 1). Thus, it is informative to analyze *price ratios* defined as defaultable discount prices relative to default-free prices for a bond with duration m : q_m/q_m^* . The price ratio of each bond is the total discounted repayment probability over the lifetime of the bond, adjusted for risk and recovery. Table 7 presents statistics for these price ratios in the model and the data. The table shows that contrary to spreads, price ratios for short-term debt are always higher than for long-term debt both in the model and in the data. Moreover, price ratios are disproportionately lower for long-term debt when spreads are high, both in the model and in the data.¹⁸

Table 7: Prices				
	DATA		MODEL	
s_S pct	q_S/q_S^*	q_L/q_L^*	q_S/q_S^*	q_L/q_L^*
< 25	0.96	0.60	0.99	0.79
< 50	0.95	0.59	0.98	0.75
\geq 50	0.85	0.43	0.85	0.57
\geq 75	0.79	0.35	0.79	0.48
Mean	0.90	0.51	0.91	0.66

The distinct dynamics of price ratios and spreads can be understood as follows. Price ratios reflect cumulative repayment probabilities adjusted for risk and recovery, whereas spreads reflect average default probabilities. Cumulative default risk for long-term debt is always larger than for short-term debt both in the data and the model. However, annualized (average) default risk can be lower on long-term debt during times when the annual default probability in the short run is larger than the annual default probability in the long run. Thus, contrary to common belief in sovereign debt markets, the interest rate spread is not a comprehensive measure of the relative cost of borrowing across different maturities of debt. In particular, in times when the probability of default is high, short-term debt may appear

¹⁸These patterns hold for price ratios in the data for Argentina, Mexico, and Russia as well.

to be more expensive than long-term debt, in the sense that it has a higher spread, although long-term debt is worse in the sense that it has a lower price, relative to the risk-free price. The connection between the dynamic behavior of prices and spreads in our model is borne out in the data as well.

The preceding discussion also indicates that the important feature of our model for generating the observed dynamics of prices and the spread curve is that the probability of default is mean-reverting: a period with high probability of default is followed by a period with lower probability of default, and vice versa. The effects of mean-reverting default probabilities on the spread curve are the same as those highlighted by Merton (1974) in the case of credit spreads for corporate debt. In our model the probability of default is endogenously mean-reverting as a result of the dynamics of the output process and debt accumulation. When output is high, it is also expected to be high in the near future, so the probability of default in the next period is low. The economy borrows a large amount at low interest rate spreads, so that in states where the economy is hit by a bad shock, default becomes more likely further in the future. In contrast, when the likelihood of imminent default is high, the economy avoids default in the next period only in states with high output. Conditional on not defaulting, then, output is expected to remain high, and the probability of default further in the future falls. The persistence and mean reversion of default and repayment probabilities driven by the dynamics of debt and income therefore rationalize the dynamic behavior of the spread curve observed in the data.

Maturity Composition

We now present the quantitative predictions for the maturity composition of debt. As discussed in Section 4, the benefits of short and long-term debt shape the dynamic behavior of the maturity composition. First, long-term bonds provide a hedge that prevents having to roll-over large amounts of debt at low prices; we find that this hedging motive is more valuable in times of high wealth. Second, short-term bonds are better at solving the lack of enforcement problem as they reduce the default likelihood and are less likely to require savings in order to be repaid. These forces give short-term debt the liquidity advantage as more resources can be transferred to the present; we find that this liquidity motive is more valuable in times of low wealth. Given the negative correlation between wealth and spreads, the borrower in our model uses long-term debt more heavily in times when spreads are low and shifts toward shorter term debt when spreads are high.

Table 8: Average Duration of New Debt Issuances

	MODEL	DATA
s_S pctl		
< 50	4.27	6.37
≥ 50	1.86	4.73
Overall	3.06	5.95

To compare issuances of long and short debt between the model and data, we compute conditional averages of the duration of new debt issuances, based on the level of the short spread. Average duration in the model is the sum of the duration (equation 18) of each new bond issuance weighted by its share in total new debt issued. Moreover, given that in the data we only compute the duration of debt issuances, and not repurchases, we do the same in the model. Table 8 reports the average duration of new debt issuances when spreads are above their median relative to when spreads are below their median in the model and in the Brazilian data. Debt duration in the model mirrors the dynamics of duration in the bond data of Brazil. In the model, average duration when spreads are low equals 4.27 years, whereas it shortens to 1.86 when spreads are high. In Brazil, the average duration of bonds issued when spreads are low equals 6.37 years and shortens to 4.73 years when spreads are high. Although the model dynamics of duration as a function of spreads mirror the data, the model underpredicts the average duration of debt and overpredicts the volatility in duration.

Table 9 provides more details about the maturity composition and the forces underlying its determination. The first row of Table 9 shows the model's portfolio – the fraction $\frac{q_S \ell_S}{q_S \ell_S + q_L \ell_L}$ of the value of new debt that is short term – conditional on different levels of the short spread. When spreads are low, the borrower issues on average 43% of debt in long-term bonds, and 57% in short-term bonds. When spreads are high, the maturity composition shifts to only 1% in long-term bonds, and 99% in short-term bonds. Moreover, the majority of the value of debt issuance is in short-term debt. As discussed in Section 4, the optimal portfolio depends on the valuations of the hedging benefits of long-term debt relative to the liquidity and cost advantage of short-term debt. In the lower section of Table 9, we reports several alternative metrics to evaluate these benefits. First, to measure the cost advantage of short-term debt, we compute the ratio of the price ratios of the two debt maturities: $\frac{q^L/q_L^*}{q^S/q_S^*}$. This ratio compares the expected discounted repayment probabilities adjusted for risk and recovery on the two types of debt. As the table shows, long-term debt is always more costly than short-term debt because it carries lower total repayment probabilities – $\frac{q^L/q_L^*}{q^S/q_S^*}$ is always less than one – which implies higher future default costs. In addition, short-term debt is disproportionately cheaper in low wealth times, as the slope of price ratios is lower, 0.67 relative to 0.77. Thus,

a larger share of short-term debt in low wealth times can be understood as a reaction to the more expensive long-term debt.

Table 9: Model Maturity Composition

	Short Spread		
	< 50 pctl	≥ 50 pctl	overall mean
$q_S \ell_S / (q_S \ell_S + q_L \ell_L)$	0.57	0.99	0.79
s_S , percent	1.2	9.7	5.5
$s_L - s_S$, percent	2.2	-0.3	1.0
$(q_L/q_L^*) / (q_S/q_S^*)$	0.77	0.67	0.72
Δ_S / Δ_L	1.43	1.67	1.46
std(value debt)	short	0.25	0.69
	long	1.08	0.03
	total	0.93	0.71
std(price debt)	short	0.09	0.21
	long	0.79	1.21
corr(value debt, y')	short	-0.45	0.83
	long	0.72	0.15
	total	0.75	0.83

A more detailed measure of the liquidity benefit of short-term debt is given by calculating the increase in consumption that would be possible by marginally increasing short-term debt, relative to the increase in consumption that is possible by issuing more long-term debt. Specifically, define $\Psi(b'_S, b'_L, b_S, b_L, y) \equiv q^S(b'_S, b'_L, y)(b'_S - \delta_S b_S) + q^L(b'_S, b'_L, y)(b'_L - \delta_L b_L)$ as the quantity of consumption that is attained with a certain debt policy b'_S, b'_L , given the state (b_S, b_L, y) . We calculate the ratio of deviations from the equilibrium debt policy for short-term debt relative to long-term debt,

$$\frac{\Delta_S}{\Delta_L} \equiv \frac{\Psi(b'_S + \varepsilon_S, b'_L, b_S, b_L, y) - \Psi(b'_S, b'_L, b_S, b_L, y)}{\Psi(b'_S, b'_L + \varepsilon_L, b_S, b_L, y) - \Psi(b'_S, b'_L, b_S, b_L, y)}$$

where ε_S and ε_L are small, and are chosen so that if bond prices were always equal to the default-free prices, the ratio would be exactly equal to 1. This ratio in our model is on average equal to 1.46. Thus, short-term debt is in equilibrium more liquid because consumption can be marginally increased more with short-term debt than with long-term debt. The reason is that price schedules for short-term debt are more lenient by having higher prices that

decrease by less as debt increases. Looking across periods, this difference is especially large in states in which spreads are high. Thus, short-term debt is particularly useful for increasing consumption when spreads are high.

Considering the individual components of this ratio, Δ_S and Δ_L , also shows why explaining the equilibrium maturity structure is more complicated than ranking short-term and long-term debt according to a “pecking order” along the lines of Myers (1984). In the context of our model, a pecking order would mean that the borrower issues one maturity of debt until its borrowing capacity is exhausted – its Δ is zero – and then issues the other maturity if additional consumption is needed. However, Δ_S , Δ_L , ℓ_S , and ℓ_L are typically all positive, meaning that the borrower issues both maturities of debt when either one alone would be able to raise consumption. The borrower does not typically borrow to the limit using either maturity, indicating that the tradeoff between short-term and long-term debt leads to an interior choice for the maturity composition.

The hedging benefits of debt can be measured by considering the variation in value of outstanding debt and future prices. The value of outstanding one period discount bonds have zero variation because in all states the value equals the face value of the bond. The value of longer-term debt varies with the state because the price depends on the varying risk-adjusted default risk. When the value of debt varies, it provides a hedge against fluctuations in the borrower’s consumption if the value and price of debt fall – meaning the present value of the debt burden declines – when income is low. Hence, we measure the hedging benefit of debt in three ways: with the standard deviation of the price q'_m and value and of each type of debt, $(1 + \delta_m q'_m) b'_m$, (and their sum) and the correlation of debt values with the borrower’s income y' .¹⁹

The table shows that on average long-term debt is a better hedge. Its price and value are more volatile than for short-term debt. The standard deviation of the long-term price and value equal 1.17 and 0.89 relative to 0.18 and 0.60 for short-term debt. Long-term debt is also more positively correlated with income with a correlation of 0.7, whereas the correlation is close to zero for short-term debt.

When spreads are high, the value of short-term debt appears more volatile and correlated with income than the value of long-term debt. This is because the quantity of long-term debt is very small relative to short-term debt – long-term debt averages only 1% of debt issuance when spreads are high. If we look at the volatility of prices, not values, during different periods, the standard deviation of the price of long-term bonds is higher than the standard

¹⁹Lustig, Sleet, and Yeltekin (2008) also measure the hedging benefits of debt with standard deviations and correlations in their work on domestic debt.

deviation of the price of short-term bonds, either within periods when spreads are high (1.21 vs. 0.21) or when spreads are low (0.79 vs. 0.09). This indicates that, even though long-term debt is still a better hedge in periods with high spreads, the liquidity benefit of short-term debt dominates the determination of the portfolio in these periods.

In summary, through the lens of our model, the maturity structure of defaultable debt in emerging markets and its covariation with spread curves and levels can be rationalized by two factors: a hedging advantage of long-term debt for insuring against fluctuations in future default risk, and a liquidity advantage of short-term debt for providing higher resources with more lenient prices.

5.3 Counterfactuals

In this section we use our calibrated model to explore how international financial market conditions affect the choice of debt maturity and default of sovereign governments. To this end, we perform two counterfactual exercises. First, we explore the role of lenders' risk aversion. We find that high risk aversion lowers default probabilities and decreases the usefulness of debt, especially of long-term debt, which leads to shorter duration of debt. Second, we eliminate recovery of debt in case of default. We find that zero recovery also lowers default probabilities and limits the usefulness of debt, but especially of short-term debt.

Risk Premia

In subsection 5.2, we assessed the contribution of risk premia by pricing bonds with the actuarially fair pricing kernel in our model, *given* the borrower's behavior. However, in our model, if the borrower were to take into account a different degree of lenders' risk aversion, debt and default decisions would change. Table 10 shows summary statistics on spreads, the maturity composition of debt, the volatility of the trade balance, and the average default probability when we re-solve our model for the case with zero and higher risk premia ($\alpha = 0$ and $\alpha = 10$ respectively) while keeping all other parameters as in the benchmark calibration.

Risk premia have large effects on spreads. As seen in the previous section, for given default probabilities risk premia increase spreads. This is the standard result in models of defaultable bond pricing models as in Duffie and Singleton (1999). However in our model with endogenous default, high risk premia can decrease the level of spreads. The reason is that risk premia make debt schedules more restricted – that is, prices fall faster with the level of debt – and the borrower issues less debt, at lower levels of spreads. High risk premia in

Table 10: Risk Premia, Spreads and Debt Maturity

s_S pct	Zero risk premia $\alpha = 0$			High risk premia $\alpha = 10$		
	s_S	s_L	duration	s_S	s_L	duration
< 50	3.1	5.5	4.7	1.3	3.4	2.0
≥ 50	13.2	12.4	1.8	11.0	13.5	1.9
Mean	8.2	10.0	3.3	6.1	8.5	1.9
std(trade balance)/std(y)	0.49			0.13		
Default probability	8.3			3.9		

our model also increase the difference between the average spread and the default probability as well as the difference between the long and short spread,

Since debt choices are endogenous in our model, risk premia also have important effects on the dynamics and maturity of debt. Higher risk premia shorten debt maturity, dampen the maturity dynamics and decrease the usefulness of debt as shown by a lower volatility of the trade balance. All debt is more expensive and used less because price schedules are very steep. Moreover, risk premia disproportionately affect the schedule of long-term debt which induces the borrower to practically never issue long-term debt. In fact the long-term spreads presented in the table are essentially shadow prices for an arbitrarily small quantity of long-term debt.²⁰

Recovery

Finally, we evaluate the impact of recovery rates on spreads and debt holdings. We resolve the model under the assumption that lenders recover zero in case of default ($\varphi(b_S, b_L) = 0$) while maintaining all other parameters as in the benchmark calibration. Table 11 contains these results.

Table 11: Recovery, Spreads and Debt Maturity

s_S pct	Benchmark			Zero Recovery		
	s_S	s_L	duration	s_S	s_L	duration
< 50	1.2	3.4	4.2	1.4	3.4	4.5
≥ 50	9.7	9.5	1.9	8.0	6.7	3.6
Mean	5.5	6.4	3.1	4.7	5.1	4.0
std(trade balance)/std(y)	0.35			0.29		
Default probability	5.5			3.8		

²⁰Broner, Lorenzoni, and Schmukler (2008) also find that higher risk premia shorten the maturity of debt when changes in risk premia are due to exogenous shocks to lenders' risk aversion.

Recovery has important effects on spreads, borrowing and default. The standard result in models of defaultable bond pricing models is that reducing recovery rates increases spreads because debt is discounted more heavily as the loss in case of default is larger for lenders. However, in our model making borrowing ex-ante more expensive means that debt schedules for short and long-term debt become more restricted which in equilibrium lowers default because the borrower issues less debt. Hence, as in the case of high risk premia, lower recovery rates makes spreads smaller and induces less usage of debt.

Table 11 shows that relative to the benchmark case (which produced an average recovery rate of 44%) having zero recovery reduces default probabilities from 5.5% to 3.8% and reduces average spreads by about 1%. In addition, zero recovery lowers the volatility of the trade balance which implies less use of debt. Lower recovery also lengthens the debt maturity by about a year and dampens the maturity dynamics. The reason why the model produces a higher mean duration when recovery is zero is that an equal recovery rate for short and long-term debt (as in the benchmark) is disproportionately beneficial for short-term debt prices, which are on average lower than the long term price. Hence, reducing recovery rates disproportionately restricts the short-term debt schedule and induces a shift in the portfolio towards long term debt. The maturity dynamics are dampened because the steep schedules restrict the liquidity and hedging benefits of short and long-term debt.

6 Conclusion

In this paper, we have developed a dynamic model to study the maturity composition of sovereign bonds. In emerging markets data, changes in the maturity composition of debt comove with changes in the term structure of spreads: when spreads on short-term debt are low, long-term spreads are higher than short-term spreads, and the maturity of debt issued is long. When short-term spreads rise, long-term spreads rise less, and the maturity of debt shortens. Our model simultaneously reproduces the patterns observed in the term structure of spreads and bond prices, and the maturity composition of debt. Changes in the spread curve, which reflects the average default probability and risk premia at different time horizons, result from the output dynamics and the endogenous dynamics of debt. Issuing long-term debt hedges future fluctuations in consumption that come from changes in debt prices. Short-term debt provides more liquidity because it allows the borrower to avoid the more severe commitment problem in repaying long-term debt. With these two forces, the model generates the pattern of issuances observed in the data. Long-term debt is issued mostly in times of high wealth and low spreads, when the hedging motive is the strongest. Short-term debt is

used more heavily in times when wealth is low and spreads are high, because expectations of the borrower's future debt and default choices restrict the availability of long-term debt more heavily than of short-term debt.

Our main innovation has been to introduce multiple, long-term assets into a quantitative, dynamic model with endogenous default. We view the resulting framework as useful for addressing a variety of other questions for which it is important to analyze a trade-off in maturity choice with defaultable debt. Natural applications are the maturity structure of consumer and corporate debt. The literature on consumer bankruptcy thus far has focused on modeling very short-term unsecured credit (Chatterjee et al. (2007) and Livshits, MacGee, and Tertilt (2007)). However, it would be interesting to analyze both long-term and short-term defaultable loans, such as mortgages and credit card debts. In addition, the mechanisms in our model are likely to be relevant in corporate debt given the similarity between our facts on emerging market spread curves and the cross section of corporate debt spread curves. Default risk has been shown to have important implications for firm dynamics (Cooley and Quadrini (2001) and Arellano, Bai, and Zhang (2007)). The model of this paper can be used to further understand how the maturity choice can influence the entry, exit, and growth of firms. Overall, our paper provides a tractable quantitative framework to study defaultable debt of multiple maturities appropriate for these questions, and has highlighted the relevant economic trade-offs important for understanding debt maturity choice in the presence of default.

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Appendix

Data Description

All the sovereign bond data are from Bloomberg. For the four countries we examine, we use all bonds with prices quoted at some point between March 1996 and May 2004, with the following exceptions. We exclude all bonds with floating-rate coupon payments, and at every date, we exclude bonds that are less than three months to maturity, following Gurkaynak, Sack, and Wright (2007). For each country, we estimate spreads starting from the first week for which at least four bond prices are available every week through the end of the sample. We use data from 110 bonds for Argentina, 71 for Brazil, 63 for Mexico, and 25 for Russia. To estimate default-free yield curves, we use data on U.S. and European government bond yields. The U.S. data are from the Federal Reserve Board, and the European data are from the European Central Bank.²¹ For constructing the quarterly maturity and duration statistics, we also include bonds issued during the sample period that did not have prices quoted, and use the estimated spread curve to construct their prices according to equation (24).

Spread Curve Estimation

We use a method proposed by Svensson (1994), and used recently by Gurkaynak, Sack, and Wright (2007) for the United States, and Broner, Lorenzoni, and Schmukler (2008) for a sample of emerging markets, to fit a spread curve to this data using a simple functional form suggested by Nelson and Siegel (1987).

A coupon bond is priced as a collection of zero-coupon bonds, each with maturity given by a coupon payment date, and face value given by the cash flow on that payment date. The price at date t of a bond issued by country i , paying an annual coupon rate c at dates n_1, n_2, \dots, n_J years into the future, is

$$p_t^i(c, \{n_j\}) = \sum_{j=1}^J \exp(-n_j r_t^i(n_j))c + \exp(-n_J r_t^i(n_J)) \quad (24)$$

with the face value of the bond paid on the last coupon date.

Spreads are defined as $s_t^i(n) = r_t^i(n) - r_t^*(n)$, where $r_t^*(n)$ is a default-free yield curve.

²¹The U.S. data are the Treasury constant maturities yields, available at <http://www.federalreserve.gov/releases/h15/data.htm>.

The European data are Euro area benchmark government bond yields, which is an average of European national government bond yields available at <http://sdw.ecb.europa.eu>.

We define spreads as a parametric function of maturity following Nelson and Siegel (1987)

$$s_t^i(n; \beta_t^i) = \beta_{1t}^i + \beta_{2t}^i \left(\frac{1 - e^{-\lambda n}}{\lambda n} \right) + \beta_{3t}^i \left(\frac{1 - e^{-\lambda n}}{\lambda n} - e^{-\lambda n} \right) \quad (25)$$

for each country i , where $\beta_t^i = (\beta_{1t}^i, \beta_{2t}^i, \beta_{3t}^i)$ and λ are parameters. For default-free bonds, we define

$$r_t^\$ (n; \beta_t) = \beta_{1t}^\$ + \beta_{2t}^\$ \left(\frac{1 - e^{-\lambda n}}{\lambda n} \right) + \beta_{3t}^\$ \left(\frac{1 - e^{-\lambda n}}{\lambda n} - e^{-\lambda n} \right) \quad (26)$$

and

$$r_t^\text{€} (n; \beta_t) = \beta_{1t}^\text{€} + \beta_{2t}^\text{€} \left(\frac{1 - e^{-\lambda n}}{\lambda n} \right) + \beta_{3t}^\text{€} \left(\frac{1 - e^{-\lambda n}}{\lambda n} - e^{-\lambda n} \right) \quad (27)$$

for US (\$) and Euro (€) bonds.

As described by Nelson and Siegel (1987) and Diebold and Li (2006), the three components of this curve correspond to a “long-term,” or “level” factor (the constant), a “short-term,” or “slope” factor (the term multiplying β_2) and a “medium-term,” or “curvature” factor (the term multiplying β_3). Linear combinations of these factors can capture a broad range of shapes for the spread curve.

We first estimate the parameters $\beta_t^\$$ and $\beta_t^\text{€}$ by OLS, using U.S. and Euro area bond yields. Throughout, we follow Diebold and Li (2006) by setting the parameter $\lambda = 0.714$, so that the term multiplying β_3 in all countries’ spread curves is maximized when $n = 2\frac{1}{2}$ years.

Then, given a set of parameters β_t^i , we use equation (24) to price each of country i ’s bonds at date t using the risk-free yield given by (26) or (27) and the spread given by (25):

$$p_t^i(c, \{n_j\}; \beta_t^i) = \sum_{j=1}^J \exp(-n_j (s_t^i(n_j; \beta_t^i) + r_t^*(n_j)))c + \exp(-n_J (s_t^i(n_J; \beta_t^i) + r_t^*(n_J))),$$

where r_t^* refers to $r_t^\$$ if the bond is denominated in U.S. dollars, or $r_t^* = r_t^\text{€}$ if the bond is denominated in a European currency.

We estimate the parameters β_t^i by nonlinear least squares to minimize the sum of squared deviations of the predicted prices, $p_t^i(c, \{n_j\}; \beta_t^i)$ from their actual values. That is, our estimated parameters solve

$$\min_{\beta_t^i} \sum (p_t^i(c, \{n_j\}; \beta_t^i) - p_t^i(c, \{n_j\}))^2,$$

where the summation is taken over all bonds issued by country i with prices available at date t . As discussed in Svensson (1994), minimizing yield to maturity errors rather than price

errors gives a better fit for short-term yields to maturity, because short-term bond prices are less sensitive to their yields to maturity than long-term bond prices.

The following features present in the data require modification of the basic bond pricing equation (24):

1. Between coupon periods, the quoted price of a bond does not include accrued interest, so we subtract from the bond price the portion of the next coupon's value that is attributed to accrued interest.
2. For bonds with principal payments guaranteed by U.S. Treasury securities, we discount the payment of principal by the risk-free yield only, without the country spread.
3. For bonds with coupon payments that increase or decrease over time with certainty ("step-up" and "step-down" bonds, respectively), we modify the sequence of payments in equation (24) accordingly.

Table 12 below displays root mean squared errors (RMSE), as a percentage of a bond's price. Errors are moderate, averaging three to seven percent of actual bond prices across the four countries. During periods when spreads are high, errors tend to be larger; for example, the RMSE of Russian bonds is 11 percent in periods when the 2-year spread is above its 90th percentile.

Table 12: Average Root Mean Squared Errors

	average RMSE (% of price)	percentile of 2-year spread					
		< 10th	< 25th	< 50th	≥ 50th	≥ 75th	≥ 90th
Argentina	5.4795	3.2470	3.3146	3.6243	7.3347	9.1872	13.2562
Brazil	6.0731	6.2179	6.1681	5.6921	6.4521	7.3479	8.2165
Mexico	2.6303	2.2326	2.2762	2.4946	2.7654	2.9852	3.2437
Russia	7.2242	4.2919	5.7055	6.4463	7.9966	9.6489	11.3806

Further Statistics on Spread Curves

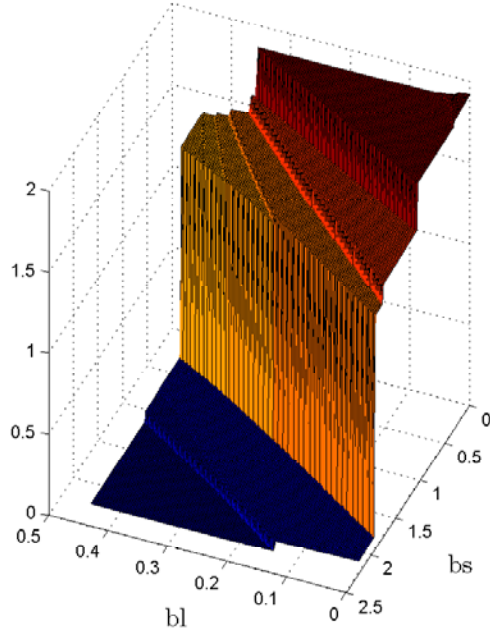
Tables 13 reports further spread curves and spread volatilities for all countries.

Table 13: Average Spreads and Volatility

	Maturity (years)	Overall (%)	Std. Dev	When 2-year spread is above/below n th percentile					
				< 10th	< 25th	< 50th	\geq 50th	\geq 75th	\geq 90th
Argentina	2	4.77	6.26	1.10	1.60	2.10	7.44	11.03	19.64
	5	5.27	3.83	1.97	2.60	3.24	7.31	9.52	14.54
	10	6.22	3.36	3.19	3.77	4.40	8.04	10.00	14.06
	15	6.63	3.30	3.70	4.25	4.87	8.39	10.31	14.12
Brazil	2	5.60	5.02	1.69	2.15	2.68	8.51	12.25	18.57
	5	6.71	4.41	3.58	3.83	4.18	9.23	12.25	17.07
	10	7.14	2.93	4.79	5.25	5.35	8.93	10.84	13.36
	15	7.28	2.41	5.24	5.81	5.81	8.75	10.19	11.74
Mexico	2	1.84	1.36	0.42	0.62	0.94	2.74	3.49	4.84
	5	2.46	0.97	1.57	1.74	2.01	2.90	3.32	4.45
	10	3.45	0.95	2.40	2.60	2.97	3.93	4.28	5.27
	15	3.87	0.98	2.72	2.93	3.35	4.40	4.74	5.66
Russia	2	4.44	2.65	1.70	2.03	2.62	6.25	7.88	9.87
	5	5.37	2.59	2.86	3.10	3.52	7.22	9.03	10.58
	10	5.10	2.25	2.86	2.81	3.39	6.79	8.11	9.36
	15	4.92	2.21	2.80	2.62	3.28	6.54	7.62	8.76

Model's Debt Price Schedules

Price Schedule for Short Debt



Price Schedule for Long Debt

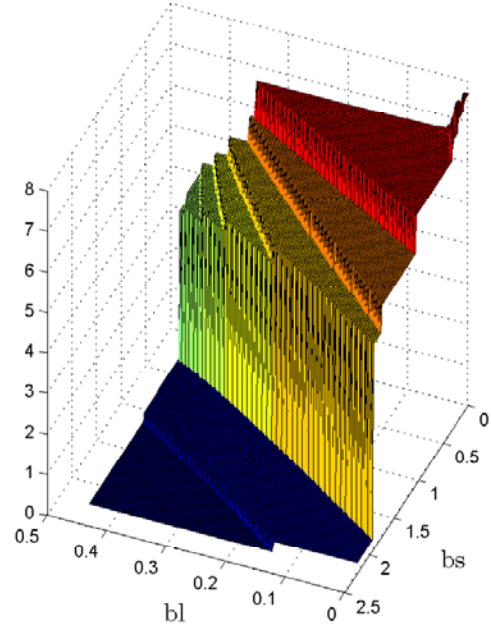


Figure 4: Price schedules for short-term and long-term debt when income is at its mean.